

Randomized dose-escalation design for drug combination cancer trials with immunotherapy

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Immunotherapy in Phase I clinical trials

Phase I conventional paradigm: the more the better

Important exceptions: molecularly targeted agents (e.g. immunotherapy).

- ★ Immune system can regulate/eliminate tumours
- ★ Low toxicity profile

Example:

immune-checkpoint proteins blocker anti-programmed-death-receptor-1 (PD1)
Pembrolizumab

- ★ None of the trials reached MTD
- ★ Plateau found. Same toxicity/activity probability for 2 and 10 mg/kg
- ★ FDA requested to focus on a lower dose level



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Combinations with immunotherapy

Immunotherapy is enough not efficacious in cancer treatment by itself

Current investigations:

- ★ the added value of immune checkpoint blockers to backbone therapy
- ★ the added value of a new drug to an immune checkpoint blocker.

One drug is administered at **full dose** while the other is **escalated**.

Objectives of the trial:

- ★ To find the maximum tolerated combination (MTC)
- ★ To detect clinically significant difference between the MTC and standard therapy alone (required by EMA)
- ★ To detect a possible dose effect in the combination



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Current approach

Current design: **one parameter CRM design** for single agent trial

Advantages:

- ★ Ability to find MTC with high probability
- ★ Well-known properties

Disadvantages:

- ★ Strong monotonicity assumption
- ★ No possibility of a plateau detection
- ★ Does not allow a statistical comparison of toxicities



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Proposals

Flexible model:

- ★ E_{max} model
- ★ A plateau in a dose-toxicity relation
- ★ Ability to model the toxicity probability on the single agent alone independently

Randomization between a control and an investigation arm

- ★ control is standard therapy
- ★ prevents a selection bias
- ★ allows statistical comparison of the toxicity
- ★ ethical



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Bayesian CRM

Combination of A (fixed) and B : $\tilde{d}_0 = \{a, b_0\}, \tilde{d}_1 = \{a, b_1\}, \dots, \tilde{d}_m = \{a, b_m\}$

Model $p_i = \psi(d_i, \theta)$; d_i is a unit-less amount of drug
 θ is a vector of parameters

Given binary outcomes, the CRM updates the posterior $f_j(\theta)$

$$f_j(\theta) = \frac{f_{j-1}(\theta)\mathcal{L}(d, y, \theta)}{\int_{\mathbb{R}^d} f_{j-1}(u)\mathcal{L}(d, y, u)du} \quad (1)$$

The posterior mean (!)

$$\hat{p}_k^{(j)} = \mathbb{E}(\psi(d_k, a) | \mathbb{Y}_j) = \int_{\mathbb{R}^d} \psi(d_k, u) f_j(u) du \quad (2)$$

Main debate: choice of model $\psi(d_k, a)$



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E_{max} model

$$\psi(d_i, E_0, E_{max}, \lambda, ED_{50}) \equiv \psi(d_i, \theta) = E_0 + \frac{d_i^\lambda E_{max}}{d_i^\lambda + ED_{50}^\lambda} \quad (3)$$

- ★ E_0 is the probability of toxicity on the control
- ★ $E_{max} + E_0$ is the maximum probability of toxicity
- ★ ED_{50} is the combination which produces $E_0 + \frac{E_{max}}{2}$
- ★ $\lambda \geq 0$ is the slope factor

Skeleton construction:

$$d_i = \hat{ED}_{50}^{(0)} \times \left(\frac{\hat{p}_i(0) - \hat{E}_0^{(0)}}{\hat{E}_{max}^{(0)} + \hat{E}_0^{(0)} - \hat{p}_i(0)} \right)^{\frac{1}{\hat{\lambda}^{(0)}}}$$

By definition, $\hat{p}_0(0) \equiv \hat{E}_0^{(0)} \rightarrow d_i = 0$



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Randomization

Assignment cohort-by-cohort

$$c = c_1 + c_2$$

- ★ c_1 be the number of patients assigned to the current best combination
- ★ c_2 be the number of patients assigned to the control, d_0 .

For instance, taking $c_1 = 3$ and $c_2 = 1$, one will end up with 25% of the total sample size being assigned to the control.

CRM with randomization results in the majority of patients on **two** combinations: control and MTC.



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Simulations setting

- ★ Sample size $n = 48$
- ★ $m = 7$ combinations
- ★ Target probability $\gamma = 0.25$;
- ★ Clinically significant difference $\tau = 0.05$
- ★ Confidence level $\alpha = 0.9$
- ★ $c_1 = 3, c_2 = 1 \rightarrow 25\%$ on the control treatment.



Characteristics

- (i) Proportion of correct recommendations
- (ii) Proportion of times the clinically significant difference is found

$$\mathcal{P} \equiv \mathbb{P}(\mathbb{P}(p_{MTC} - p_{control} \geq \tau) > \alpha) \quad (4)$$

- (iii) Goodness of fit measure

$$NMSE = \frac{1}{N} \sum_{j=1}^N \sqrt{\frac{\sum_{i=1}^n (p_i - \psi(d_i, \hat{\theta}^{(j)}))^2}{\sum_{i=1}^n (p_i - \hat{p}_i^{opt})^2}} \quad (5)$$



Prior and comparators

Skeleton

$$\mathbb{P}_0 = [\mathbf{0.08}, 0.25, 0.35, 0.45, 0.55, 0.65, 0.70, 0.75]^T$$

Information to construct prior distributions for model parameters:

- (i) Control: upper bound of the 95% credibility interval is 0.25.
- (ii) Prior MTC: upper bound of the 95% credibility interval is 0.80.

$$E_0 \sim \mathbb{B}(0.8, 10-0.8); \lambda \sim \Gamma(1, 1); E_{max}|E_0 \sim \mathbb{U}[0, 1-E_0]; ED_{50} \sim \Gamma(0.4, 0.4);$$



Comparators

(P1) **One-parameter** power model (**no randomization**):

$$\psi(d_i, z) = d_i^z.$$

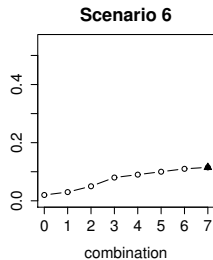
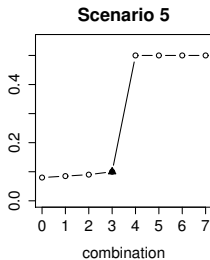
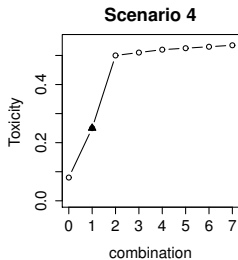
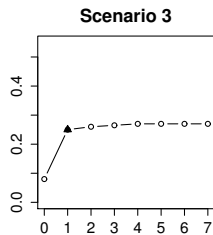
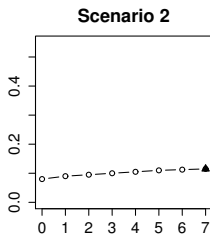
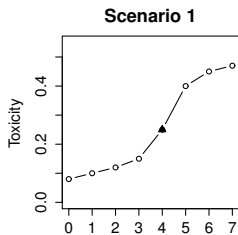
(L2) **Two-parameter** logistic models

$$\psi(d_i, \beta_1, \beta_2) = \frac{\exp(\log(\beta_1) + \beta_2 d_i)}{1 + \exp(\log(\beta_1) + \beta_2 d_i)}$$

with (R) and without randomization.



Scenarios



Results (I)

	d_0	d_1	d_2	d_3	d_4	d_5	d_6	d_7	TR
Sc 1	0.08	0.10	0.12	0.15	0.25	0.40	0.45	0.47	
$E_{max}(R)$	0.0	2.1	8.7	22.9	45.1	13.6	3.8	3.7	19.1
$L2(R)$	0.0	1.8	8.3	26.5	44.6	13.8	3.1	2.0	19.2
$L2$	0.4	0.1	3.3	24.9	52.6	16.4	2.1	1.3	25.1
$P1$	0.0	1.0	4.2	16.5	51.4	20.4	5.5	1.0	29.2
Sc 2	0.08	0.09	0.095	0.10	0.10	0.11	0.11	0.11	
$E_{max}(R)$	0.0	0.3	0.5	0.9	0.5	1.6	1.9	94.4	10.2
$L2(R)$	0.0	0.62	0.75	1.50	3.00	2.13	1.75	90.2	10.6
$L2$	0.3	0.1	0.3	0.5	1.8	1.0	0.5	95.6	11.4
$P1$	0.0	0.0	0.0	0.0	0.0	0.6	0.8	98.7	11.3



Results (II)

	d_0	d_1	d_2	d_3	d_4	d_5	d_6	d_7	TR
Scenario 4	0.08	0.25	0.50	0.51	0.52	0.52	0.53	0.54	
$E_{max}(R)$	5.1	82.1	11.5	0.9	0.3	0.1	0.0	0.0	24.4
$L2(R)$	4.1	86.4	8.4	0.9	0.3	0.0	0.0	0.0	24.6
$L2$	20.9	75.0	4.0	0.1	0.0	0.0	0.0	0.0	26.9
$P1$	20.2	71.7	7.4	0.7	0.0	0.0	0.0	0.0	31.0
Scenario 5	0.08	0.09	0.09	0.10	0.5	0.5	0.5	0.5	
$E_{max}(R)$	0.0	1.9	7.7	57.1	31.2	1.6	0.1	0.4	18.8
$L2(R)$	0.0	1.5	9.5	56.4	30.2	1.4	0.8	0.3	18.2
$L2$	0.1	0.1	1.4	62.7	35.5	0.2	0.1	0.0	25.1
$P1$	0.0	0.0	0.0	54.9	36.4	7.0	1.2	0.5	33.2



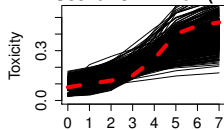
Results (III)

		Sc 1	Sc 2	Sc 3	Sc 4	Sc 5	Sc 6
$E_{(R)}$	\mathcal{P}	74.7%	16.3%	66.7%	71.5%	67.9%	24.2%
	NMSE	1.7	2.0	2.2	3.5	3.1	1.5
$L2_{(R)}$	\mathcal{P}	71.8%	14.4%	59.7%	64.1%	74.9%	20.9%
	NMSE	1.9	2.2	7.2	4.4	3.4	1.5
$L2$	\mathcal{P}	61.5%	15.7%	50.1%	50.2%	64.7 %	18.1%
	NMSE	2.0	2.5	7.7	5.8	3.6	1.6
$P1$	\mathcal{P}	99.9%	93.7%	99.7%	99.7%	99.4%	95.3%
	NMSE	2.1	2.6	7.8	6.1	3.8	2.1

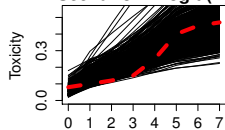


Fitted curves

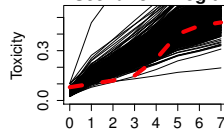
Scenario 1. Emax (R)



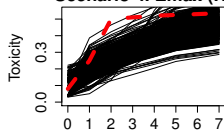
Scenario 1. Logit (R)



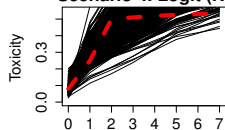
Scenario 1. Logit



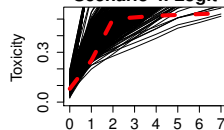
Scenario 4. Emax (R)



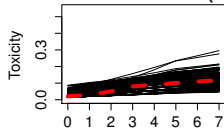
Scenario 4. Logit (R)



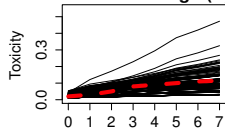
Scenario 4. Logit



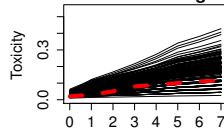
Scenario 6. Emax (R)



Scenario 6. Logit (R)



Scenario 6. Logit



combination

combination

combination



Sensitivity analysis

Prior distributions:

$$E_0 \sim \mathbb{B}(0, 8, 10-0.8), \quad \lambda \sim \Gamma(c_1, c_2), \quad E_{max}|E_0 \sim \mathbb{U}[0, 1-E_0], \quad ED_{50} \sim \Gamma(c_3, c_4).$$

- ★ Recommendation: An informative for λ and an uninformative for ED_{50}

Randomization proportion

- ★ Recommendation: 20%-25% on the control arm



Conclusions

- ★ Randomization and E_{max} model allow to identify clinically significant differences **with higher probability** than alternatives.
- ★ The cost of randomization: **a small reduction** in the proportion of correct recommendations in some scenarios.
- ★ The randomization helps to **overcome problems with fitting**
- ★ This design should be considered if **not only MTC identification** is of interest



Further work

- ★ Large variance of number of patients on the MTC.
- ★ Further investigation on the fitting problem
- ★ Phase II trials: a statistical comparison of the optimal combination and control effectivenesses

